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Modelling of economic growth determinants in Ghana in the presence of structural breaks

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Abstract

The study investigates the determinants of economic growth in the presence of structural breaks using the Gregory and Hansen co-integration model in the Ghanaian economy for the period 1971 to 2011. The variables considered as the determinants were financial development, inflation, government expenditure, and trade openness. The empirical findings are in agreement with the existence of co-integration in the presence of structural breaks. The study shows that there are structural breaks that coincide with identified climatic, economic, and political shocks. The finding does not support short-run nexus between growth and the determinants considered in the study. However, financial development, government expenditures, and trade openness are the long-run determinants of growth. In respect of policy, government-initiated structural reforms aimed at ensuring growth is of limited value, since the effect of such reforms on the long-run growth path will be offset by other shocks to the economy. Besides, in other to achieve sustainable economic growth, policymakers should put in place strategies to ensure that the financial sector is properly strengthening, trade is appropriately liberalized and government expenditure is targeted at the productive sector of the economy. Future studies, in line with the focus of the current study, based on panel cointegration, accounting for structural beaks effect, is worth embarking on.

Keywords: Co-integration, Long run, Short run, financial development, trade openness JEL Classification: F1, F41, F43, G1, G2, H1

1. Introduction

For many years, the determinants of economic growth have attracted the attention of both policymakers and research in various fields such as applied economics, finance, econometrics, development economics and many others. This sustained interest in research in this area stems from the fact that there are many benefits nations developed from economic growth.

There has been many policies and theoretical focused empirical research in the area of longrun and causality link between economic growth and its determinants, using various estimation methods such as Johansen (1988) and Johansen and Juselius (1990), Autoregressive distributed lag models (Pesaran et al., 2001), Engle-Granger test (1987), Gregory (1996) and Hansen test (1992), and single equation Ordinary Least Energy Square.

The focus of the current research is on the determinants of economic growth by examining the role of financial development, government expenditure, trade openness, and inflation in influencing economic growth in both long run and short run. A review of related works revealed that many prior studies did not consider the effect of structural breaks in determining the determinants of economic growth (Omisakin and Adeniyi, 2014).

The present research adds to the existing literature by investigating the determinants of economic growth by accounting for the effect of endogenous structural breaks in the cointegration analysis. Besides, the paper makes a contribution to the literature on unit roots by a single structural break and a multiple break-in time series analysis.

The research contributes to both exogenous and endogenous growth theories by providing better understanding empirically of the factors influencing the growth of an economy and the channels of effect. These factors are inflation, government expenditures, and trade openness.

Another meaningful contribution to the growth theories is that it also establishes the role of variables such as finance in growth, which has been neglected in the earlier models as indicated by Stern (2000) and Sadorsky (2010).

The findings, also, are expected to help policies makers on the economy to plan properly in the area of economic growth. This is so since any economic growth projections in emerging economies such as Ghana without considering financial development, trade openness, government expenditures as an explanatory variable might provide an inaccurate estimate of actual economic growth and unduly interfere with the polices.

The research is based on research questions such as; what is the effect of structural change on economic growth? What are the short-run and long-run determinants of economic growth in the presence of structural breaks? Which explanatory variables can predict economic growth? The following hypothesis was considered in the study. There are statistically significant structural breaks that coincide with identified climatic, economic, and political shocks. The explanatory variables in the model estimated to have a statistically significant effect on economic growth.

The next sections of the paper are organised into five parts. The review of related works (2); methodology section (section 3); results section (section 4), discussion section (section 5); and conclusion section (6).

2. Review of Related Works

The review of related works is organised into four main parts. They are financial development and economic growth; inflation and economic growth; government expenditure and economic growth; and trade openness and economic growth.

2.1 Finance and Growth

Theoretically and empirically, the link between finance and growth have been examined. The theoretical findings are reported in the works of prior researchers such as Bagehot (1873); Schumpeter (1911); Gurley and Shaw (1955); Goldsmith (1969); McKinnon (1973); According to these authors, economic growth is a function of financial development in both advanced and early stages of development of any nation. Scarce resources from surplus to deficit side of an economy are channelled through an efficient financial system for growth and development.

The findings of the empirical works are reported in the works of researchers such as Goldsmith (1969); Greene and Villanueva (1991); Demetriades and Devereux (1992); Alesina, Grilli and Milesi-Ferretti (1993); King and Levine (1993); Grabel (1995); Demetriades and Hussein (1996); Demetriades, et al. (1996); Eatwell (1996); Hermes (1996); Arestis and Demetriades (1997); Levine and Zervos (1998); Rajan and Zingales (1998); Rodrik (1998); Kang and Sawada (2000); Rivera-Batiz (2001); Levine (2001); Shan, Morris and Sun (2001); Kar and Pentecos (2002); Calderón (2003); Prasad, Rogoff, Wei and Köse (2003); Arestis (2004, 2005); Dritsakis and Adamopoulos (2004); Levine (2004); Bonfiglioli (2005); Galindo, Schiantarelli, and Weiss (2005); Klein (2005); Mansor (2005); Liang (2006); Loayza and Ranciere (2006); Ranciere, Tornell and Westermann (2006); Naceur (2007); Shrestha, and Chowdhury (2007); Yapraklı (2007); Ang (2009); Lee and Shin (2008); Luintel et al. (2008); Klein and Olivei (2008); Bick, Kremer and Nautz (2009); Hepsağ (2009); Mohammad et al. (2009); Sergii (2009); Zheng and Yu (2009); Adamopoulos (2010); Akinlo and Egbetunde (2010); Dabos and Gantman (2010); Odeniran and Udeaja (2010); Ahmed and Suliman (2011); Nouri and Samimi (2011); Tabi and Ondoa (2011); Baliamoune-Lutz (2013);

Kapingura (2013); Samargandi et al. (2013); Balago (2014); Fang and Jiang (2014); Adeniyi et al. (2015); Dinar, Dalgiç, and İyidoğan (2015); Dilek (2016); Qamarzumman (2017).

The findings of these empirical works indicate various links between financial development and economic growth. They are positive, negative, direct, indirect, growth-driven finance hypothesis, finance-driven growth hypothesis, bidirectional hypothesis, and neutral nexus between finance and economic growth.

Olusegun and Oluwatosin (2014) examined the link between finance and economic growth for the ECOWAS countries for the period 1970 to 2008 using yearly time series data. They reported significant cointegration in the presence of structural breaks. Their results indicate that government expenditure, trade openness, capital investment enhanced economic growth. The Granger causality test did not support growth-driven finance assumption.

In a study in which Dinar, Dalgiç, and İyidoğan (2015) used time-series data on Turkey for the period 1998 to 2012. They obtained results indicating a stable long-run link between finance and economic growth in the presence of structural breaks, with causality running to finance from growth. The study was based on Gregory and Hansen cointegration analysis and Zivot and Andrews stationarity test.

Elijah and Hamza (2019) investigated the nexus between finance and economic growth in the presence of structural breaks, for Nigeria with trade openness as a control variable, using yearly time series data for the period 1981 to 2015. They obtained results indicating a stable long-run relationship between finance and economic growth, with a negative link.

2.2 Inflation-Growth

Many theories account for the nexus between inflation and economic growth in the growth literature. Some of the theorists (see Swan, 1956; Solow, 1956; Mundell, 1963; Tobin, 1965; Sidrauski, 1967; Stockman, 1981; Blanchard & Kiyotaki, 1987; Greenwood & Huffman, 1987; Cooley and Hansen, 1989; Tobin, 1972; Haslag, 1995; Marquis & Reffert, 1995; Dornbusch, et al., 1996; Haslag, 1997; Todaro, 2000; Gillman, Harris & Matyas, 2004) posit that inflation is positively linked to growth, negatively linked to growth and neutrally linked to growth.

The theories are Classical, Keynesian, Neo-Keynesian, Monetarist, Neo-classical and Endogenous growth theories. For example, in the classical growth model inflation is related to growth. The supply-side focuses on the effect of savings and investments in an economy and that labour, land and capital are important in the growth model. The aggregate demand (AD) and aggregate supply (AS) model is used to explain the inflation-growth nexus in the Keynesian and Neo-Keynesian model. In the monetarism model, the quantity theory of money is used to explain the causes of inflation as resulting from monetary growth. Inflation affects growth through capital accumulation and investment in the economy by the endogenous growth model and neo-classical growth model.

A lot of empirical studies on the inflation-growth nexus in all economics testing various hypothesis such positive link, negative link, neutral link, inflation-driven growth, growth-driven inflation, as well as a bidirectional link exist. Some studies have also focused on the level of inflation and growth link, short-run link and long-run link. The findings have been mixed in the empirical literature.

The findings are found in the works of various researchers (Fischer, 1993; Ghosh & Phillips, 1998; Shitundu & Luvanda, 2000; Mallik & Chowdhury, 2001; Gokal & Hanif, 2004; Mubarik, 2005; Bick, Kremer and Nautz, 2009; Mohammad et al., 2009; Munir, Mansur & Furuoka, 2009; Sergii, 2009; Frimpong & Oteng-Abayie, 2010; Espinoza et al., 2010; Hasanov, 2011; Marbuah, 2010; Ahmed & Suliman, 2011; Kasidi & Mwakanemela, 2013; Marbuah, 2013; Thanh, 2015; Van Eyden et al., 2015; Ibarra & Trupkin, 2016; Ndoricimpa, 2017).

Some studies found a negative effect on growth in a linear model (Stockman, 1981; Kormendi & Meguire, 1985; De Gregorio,1993; Fischer, 1993; Barro, 1995, 1996; Dewan et al., 1999; Dewan & Hussein, 2001; Gokal & Hanif, 2004; Munir, Mansur & Furuoka, 2009). Other studies (Levine & Zervos, 1993; Sala-i-Martin, 1997) indicate inflation does not influence growth especially in the presence of other control variables.

Studies (Fischer, 1993; Barro, 1995, 1996; Sarel, 1996; Andres & Hernando, 1997; Bruno & Easterly, 1998; Ghosh & Phillips, 1998; Khan & Senhadji, 2001; Caner & Hansen, 2004; Sweidan, 2004; Drukker et al., 2005; Hodge; 2005; Mortaza, 2005; Mubarik, 2005; Fabayo and Ajilore, 2006; Kremer et al., 2009; Bick, 2010; Espinoza et al., 2010; Frimpong & Oteng-Abayie, 2010; Omay & Kan, 2010; Phiri, 2010; Quartey, 2010; Salami & Kelikume; 2010; Lopez-Villavicencio & Mignon, 2011; Mohanty, Chakraborty, Das & Jogn, 2011; Ajideand Olukemi, 2012; Kremer et al., 2013; Seleteng et al., 2013 Vinayagathasan, 2013; Eggoh & Muhammad, 2014; Thanh, 2015; Van Eyden et al., 2015; Ibarra & Trupkin, 2016; Ndoricimpa, 2017) that have investigated the level of inflation needed to influence growth and the nonlinear link between inflation and growth conclude that not all levels of inflation are relevant influencing growth and that the link between inflation and growth is not linear.

In the study by Dewan and Hussein (2001) reported a significant negative nexus between growth and inflation for 41 middle-income countries. In a similar study, Faria and Carneiro (2001) on the economy of Brazil using yearly data between 1980 and 1995, they reported of a significant negative link between inflation and growth in the short run.

In a Jordan study by Sweidan (2004) for the period 1970 to 2003 positive link between growth and inflation was established below a threshold of 2 and a significant negative nexus at a threshold of 2.

In a Taiwan study, Lee and Wong (2005) for the period 1965 to 2002 assessed the inflationgrowth link and indicated that inflation is negatively related to growth when the threshold exceeds 7.3%. They also investigated the nexus for the Japan economy for the period 1970 to 2001 and reported of threshold and concluded that beyond threshold levels of 2.52% and 9.66% inflation is not favourable to growth.

For the period 1976 to 2007, Bhaduri (2007) in a study of the Indian economy indicated that inflation and growth are negatively related in the short run and not in the long run.

In Turkey, for the period 1987-2006, Erbaykal and Okuyan (2008) investigated the growthinflation link and reported significant short-term link and no long-term link between inflation and growth. In a Malaysian study by Munir et al. (2009) for the period 1970 to 2005 in a threshold analysis reported of a threshold effect and indicated that at higher levels of the inflation rate, growth is affected negatively.

In a Ghanaian study, for the period 1960 to 2008, Frimpong and Oteng-Abayie (2010) indicated a negative link between inflation and growth in the presence of breaks. Threshold levels were identified in their study. In a similar Ghanaian study, for the period 1970 to 2006, Quartey (2010) examined the effect of inflation on growth in the presence of breaks established threshold effect and reported of a significant negative link between growth and inflation.

Marbuah (2011) further examined the growth-inflation link for Ghana for the period 1955 to 2009 in the presence of structural breaks and concluded that higher inflation rates affect growth negatively and support the policy of inflation targeting in the Ghanaian economy. In a similar threshold analysis for the period 1996 to 1997 and 2010 to 2011 using quarterly data for India, Mohanty et al. (2011) reported of a significant positive link between inflation rate and growth in the presence of structural breaks and concluded that lower levels of inflation are appropriate for economic growth.

Seleteng et al. (2013) examined the inflation-growth link for SADC countries accounting for the threshold effect and concluded that a higher inflation rate negatively influences growth for the period under study. Kremer et al. (2013) in a panel study for 124 countries examined the effect of inflation on growth in a threshold analysis and concluded that lower inflation rates enhance growth and higher rates impede growth. For 32 countries in Asia, Vinayagathasan (2013) investigated the growth-inflation link accounting for threshold effect reported that lower inflation rates (above 5%) enhance growth whereas higher rates negatively affect growth.

In a panel study on the link between inflation and growth accounting for thresholds effect, for developed and developing economies, Eggoh and Muhammad (2014) study reported that higher inflation rates negatively affect growth.

In a study on ASEAN countries, Thanh (2015) examined the growth-inflation link accounting for the effect of a threshold, concluded that higher inflation rates are not good for economic growth. Van Eyden et al. (2015) in a further study on growth-inflation considering the effect of a threshold, reported a negative effect of inflation on growth at higher levels of inflation rates.

In 138 country study considering the effect of threshold on the growth-inflation link, Ibarra and Trupkin (2016) conclude that inflation negatively affects growth at a higher inflation rate and that the threshold rate for nonindustrial countries is about 19% and that for industrial countries is about 5%.

In an India study, Kallah (2018) reported of a significant positive link between inflation and growth in the presence of breaks at different threshold levels and concluded that lower levels of inflation rates are important for economic growth in India.

2.3 Government Expenditure-Growth

Two main theorists (Wagner and Keynes) explain the effect of government expenditure on economic growth in the growth literature. These two theorists give opposing explanations on the effect of government expenditure on growth, hence the empirical validation of these two approaches have yielded mixed findings and conclusions in the literature (Grossman, 1988).

According to the Wagner (1883) approach, government expenditure is an endogenous variable in the growth model (Musgrave and Musgrave, 1988; Cooray, 2009) whereas, in the Keynes (1936) model, government expenditure is exogenous variable (Abdullah, 2000).

The empirical findings on the link between government-growth link are reported in the works of researchers (Barro, 1991; Ghali, 1998; Bajo-Rudio, 2000; Albatel, 2002; Abu-Bader & Abu-Qarn, 2003; Niloy et al., 2003; Pevcin, 2003; Akpan, 2005; Mitchell, 2005; Gregoriou & Ghosh, 2007; Olugbenga & Owoye, 2007; Afonso & Furceri, 2008; Liu et al., 2008; Ranjan & Sharma, 2008; Sharma & Ramful, 2008; Alexiou, 2009; Cooray, 2009; Frimpong & Oteng-Agbaiye, 2009; Kumar, 2009; Maku, 2009; Mohammad et al., 2009; Pham, 2009; Bergh & Karlsson, 2010; Ighodaro & Oriakhi, 2010; Nurudeen & Usman, 2010; Taban, 2010; Verma and Arora, 2010; Adeniyi & Bashir, 2011; Afonso & Furceri, 2010; Afonso & Jalles, 2011; Usman et al., 2011; Adewara & Oloni, 2012; Boroujli, Amin, Mehrara, & Abrishami, 2013; Oyinlola & Akinnibosun, 2013; Srinivasan 2013).

The main findings as reported in the literature are that there is mixed findings, neutral findings, negative findings and positive findings. For example, Sharma and Ramful (2008) in accounting for the effect of structural breaks in Australia and the USA study, found a significant long-run link between expenditure and growth and no significant nexus in the model structural breaks were not accounted for.

Kumar (2009) studied the link between growth and government spending some East Asian economies considering the effect of structural breaks during the period 1960 to 2007. The Gregory and Hansen model used indicated a significant long-run link between growth and expenditures. Wagner's hypothesis was supported in the countries in the study aside in Hong Kong.

Oyinlola and Akinnibosun, (2013) for the period 1970 to 2009 investigated the growthexpenditure link for Nigerian economy accounting for structural breaks supported Wagner's hypothesis and concluded that in the long-run growth and government expenditure are related and that economic growth is the object of government expenditure.

For the post-apartheid period 1994 to 2015 in South Africa, Mlilo and Netshikulwe (2017) examined the growth-expenditure nexus accounting for the effect of structural breaks. There was no significant evidence of a stable long-run link between growth and expenditure. Wagner's hypothesis was rejected in support of the Keynesian hypothesis for the period under investigation.

2.4 <u>Trade-Growth</u>

The trade-growth nexus has received attention in the growth literature theoretically and empirically with contradicting findings though. The theoretical concepts are the exogenous growth model and the new growth model. According to the exogenous growth, model trade does not influence long-run growth (Rivera-Batiz & Romer 1991). In the context of the new growth model, trade influence long-run growth through resource allocation and world integration. This leads to growth in the developing economies (Grossman & Helpman 1991b).

Various empirical studies have examined the link between trade and growth in different kinds of hypothesis such as trade-driven growth hypothesis, growth-driven trade hypothesis, bidirectional causality, positive link, negative link, and neutral link. The empirical verification of these hypothesis has produced mixed findings in the growth literature. The empirical findings are found in the works of various researcher such as Karacaer and Kapusuzoglu (2010); Ozturk and Acaravci (2010); Manni and Afzal (2012); Dawson and Sanjuán-López (2013); Belloumi (2014); Soliu and Ibrahim (2014); Brana (2016); Iamsiraroj (2016); Uslu, Aydoğan, and Ketenci (2015); Alsamara, Mrabet, Barkat, and Elafif (2019).

Some of the studies did account for the effect of structural breaks such as the study by Ozturk and Acaravci (2010) for Turkey which reported of no significant long-run relationship between trade (proxied by export) in the presence of foreign direct investment, and a stable significant long-run link between trade (proxied by import) in the presence of foreign direct investment.

Dawson and Sanjuán-López (2013) investigated the trade-growth nexus for 47 developing countries during the period 1970 to 2004 in a panel model that accounted for the effect of structural breaks. The findings indicated a significant long-run link between trade and growth with bidirectional causality between the two variables in a bivariate model.

Uslu, Aydoğan, and Ketenci (2015) studied the impact of trade on growth in 21 developing countries in a panel model for the period 1995 to 2013 using quarterly data considering the influence of structural breaks. Their findings showed that trade account for growth and the effect of structural breaks is significant in the sense that the effect decrease in the presence of structural breaks.

Alsamara, Mrabet, Barkat, and Elafif (2019) examined the effect of trade on Turkey's economic growth for the period 1960-2014 period accounting for the effect of structural breaks. They reported a significant positive effect of trade on growth.

Nketiah, Cai, Adjei, and Boamah (2020) investigated the effect of trade on growth in the presence of foreign direct investment without accounting for the effect of structural breaks and reported that growth in Ghana for the period 1975 to 2017 is a function of trade openness.

3. Methodology

The paper aims to examine the determinants of economic growth in the presence of structural breaks. This is achieved by examining the unit root properties using Zivot-Andrews (1992) and Clemente, Montanes and Reyes (1998) tests. After that, the long-run relationship among the growth and the determinants are estimated using the Gregory and Hansen Methodology with Structural breaks (Sweidan, 2004).

3.1 <u>Stationarity Test</u>

The Zivot-Andrews (1992) and Clemente, Montanes and Reyes (1998) tests are used in considering structural breaks in time series analysis. The Zivot-Andrews (1992) unit root test (ZA) allows for the examination and investigation of unit root to be examined with one endogenously determined structural break. The ZA unit root test estimates the break date and does not treat the break date as fixed. A single break is allowed in the intercept and the trend of the series variables under investigation. Different dummy variables are used for each break date in the model. The ZA test is as specified in equations (1), (2) and (3) following the works of Perron's ADF test.

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$$\begin{aligned} x_{t} &= \hat{\mu}^{A} + \hat{\theta}^{A} du_{t}(\hat{\lambda}) + \hat{\beta}^{A} t + \hat{\alpha}^{A} x_{t-1} + \sum_{j=1}^{L} \hat{c}_{j}^{A} \Delta x_{t-j} \\ &+ \hat{e}_{t}, \dots \dots (1) \\ x_{t} &= \hat{\mu}^{B} + \hat{\beta}^{B} t + \hat{\gamma}^{B} dT^{*}{}_{t}(\hat{\lambda}) + \hat{\alpha}^{B} x_{t-1} + \sum_{j=1}^{K} \hat{c}_{j}^{B} \Delta x_{t-j} + \hat{e}_{t}, \dots \dots (2) \\ x_{t} &= \hat{\mu}^{C} + \hat{\theta}^{C} du_{t}(\hat{\lambda}) + \hat{\beta}^{C} t + \hat{\gamma}^{C} dT^{*}{}_{t}(\hat{\lambda}) + \hat{\alpha}^{C} x_{t-1} + \sum_{j=1}^{K} \hat{c}_{j}^{C} \Delta x_{t-j} + \hat{e}_{t}, \dots (3) \end{aligned}$$

If du_t (λ) =1, t>T λ , 0 otherwise; dT^{*}_t(λ) = t-T λ if t>T λ , 0 otherwise. In equations (1), (2), and (3) the estimated values of the break fraction are the lamda. Equation four (4) specifies the null hypothesis which states that the series under investigation are integrated with no exogenous structural break. The alternative hypothesis indicates that the series variables under investigation are a trend-stationary process with a breakpoint that occurs in the trend at a point that is not known. In this case, the breakpoint is estimated at a point that gives the most weight to the alternative assumption.

Information criteria such as Schwartz Information Criterion (SIC); Akaike information criterion (AIC); Bayesian information criterion (BIC) and Hannan-Quinn information criterion (HQI) are used for the lag length selection. The t-test is used also in the lag selection process.

Empirical studies indicate that some series of variables exhibit multiple breaks and not just one break (Perman & Byrne, 2006). Clemente, Montanes, and Reyes (1998) developed a unit root test that accounts for multiple structural breaks based on the work of Perron and Vogelsang (1992). Equations (5) and (6) specifies the null and alternative hypothesis respectively.

The dummy variable in the equations is dTb_{it} which is =1 if $t=Tb_i+1,0$ otherwise. $du_{it} = 1$ If t>Tb_i, $du_{it} = 1$ (for i=1, 2), 0 otherwise. Tb1 and Tb2 stand for the period for the breakpoint. Clemente et al. (1998) indicated that Tbi = λiT (for i=1, 2) for the range $0 < \lambda_i < 1$ and $\lambda_2 > \lambda_1$. The breakpoints are estimated based on innovative outliers and additive outliers. The innovative outlier model is as specified in equation (7).

$$y_{t} = \mu + \rho y_{t-1} + \delta_{1} dT b_{1t} + \delta_{2} dT b_{2t} + \theta_{1} du_{1t} + \theta_{2} du_{2t} + \sum_{i=1}^{\kappa} c_{j} \Delta y_{t-i} + \varepsilon_{t}, \dots \dots (7)$$

The asymptotic distribution of the test statistic for the model is given based on the assumption in the model (8)

$$0 < \lambda_0 < \lambda_1, \lambda_2 < 1 - \lambda_0 < 1, \dots (8)$$

In equation (8), λ_1 and λ_2 take the values of [(t+2)/T, (T-1/T)]. To avoid situations where breaks occur in conservative time period Clemente et al. (1998) assumed that $\lambda_2 > \lambda_1 + 1$. In the case of additive outliers, the models are specified in equations (9) and (10).

$$x_t = \mu + \theta_1 du_{1t} + \theta_2 du_{2t} + \tilde{x}....(9)$$

In equation (9) the deterministic part of the model is eliminated according to Clemente et al. (1998) in estimation. Equation (10) is estimated based on an assumption that $\rho = 1$ after the

estimation of equation (9). The null assumption that $\rho = 1$ is tested using the minimum values of the t-ratio.

The dummy variable is dTb_{i} , and it is included in the model to ensure that min $t^{A0}_{\ \varrho^{\wedge}}(\lambda_1, \lambda_2)$ converges to the distribution as stated by Clemente et al. (1998). Thus,

$$\min t_{\rho^{\Lambda^{A0}}}(\lambda_1,\lambda_2) \to \inf \lambda = \Lambda \frac{H}{[\lambda_1(\lambda_2 - \lambda_1)(1 - \lambda_2)]^{1/2}k^{1/2}}\dots\dots\dots\dots(11)$$

3.2 Gregory and Hansen Cointegration Test

The paper employs the Gregory and Hansen model of cointegration to account for the effect of endogenous structural breaks in investigating the long run link. Equations (12), (13), (14), and (15) are specified by Gregory and Hansen (1996a; 1996b) to test for cointegration link on the null hypothesis no cointegration against the alternative hypothesis, with two variables, dependent variable (Y) and explanatory variable (X). In the equations, error terms are = e; t= time subscript; k= break date.

Model A: Level Shift

Model B: Level Shift with Trend

Model C: Regime Shift with a change in Intercept and Slope coefficients

$$Y_t = \theta_1 + \theta_2 f_{tk} + \delta_1 t + a_1 X_t + a_2 X_t f_{tk} + e_t, \dots, (14)$$

Model D: Regime Shift with change Intercept, Slope coefficients and Trend

$$Y_t = \theta_1 + \theta_2 f_{tk} + \delta_1 t + \delta_2 t f_{tk} + a_1 X_t + a_2 X_t f_{tk} + e_t, \dots, (15)$$

In estimating dummy variables in equations (12) to (15), Gregory and Hansen (1996) specified equation (16).

 $f_{ik} = 0 \text{ if } t \le k \text{ and } f_{ik} = 1 \text{ if } t \ge k$ (16)

In the equations, the break date is determined through the estimation of the cointegration equations for all possible break dates in the series in the study (Gregory & Hansen,1996),). The break date is selected using the minimum t-statistics or the maximum absolute values of the ADF test statistics. The works of MacKinnon (1991) was used by Gregory and Hansen (1996) to develop the critical values in Engle-Granger model which is use to examine cointegration nexus accounting for unknown structural breaks. Lee and Chang (2005) indicate that the presence of structural breaks distorts empirical results when not considered in analysis.

3.3 <u>The Conceptual Framework of the Empirical Model</u>

A multivariate (multiple regression) models are used to examine the determinants of economic growth with growth as the dependent variable and inflation, trade openness, government expenditure, and financial development as independent variables. The model is conceptualized as in equation (17), where M stands for the dependent variable and N, the independent variables.

$$M_t = \beta_0 + \beta_1 N_1 + \ldots + \beta_p N_p + \varepsilon_t \dots \dots \dots \dots \dots (17)$$

3.4 <u>Data</u>

The paper is based on annual data obtained from the World Bank database (World Development Indicator-WDI) were employed for the analysis. The study period is from 1971 to 2011.

Source	Data Description
WDI	Economic growth (y) proxied by gross domestic product
WDI	Government expenditure (GE)
WDI	Inflation (IN)
WDI	Trade openness (OPEN)
WDI	Financial development (M2) proxied by money supply (M2 monetary aggregate)
	Trade openness (OPEN)

Table 1 Data Description

4. Empirical Results

4.1 Zivot-Andrews Unit root Test Results

The Zivot-Andrews test was used to test for unit root allowing for an endogenously determined structural break. The results are reported in Table 2. The test is based on the null hypothesis of unit root against the alternative hypothesis of no unit root. The null hypothesis of unit root cannot be rejected. The series variables are unit root with structural breaks. The break dates coincide with known national dates in the Ghanaian economy. For example, in 1983, there was a drought in Ghana, which affects these variables.

Series (Level)	t-statistic	Optimal Breakpoints	Decisions
OPEN	-2.772	2005	Unit root
У	0.401	2005	Unit root
GE	-3.473	1992	Unit root
M2	-3.596	1979	Unit root
IN	2.141	2003	Unit root

Table 2 ZA (1992) unit root tests Results

(Author's computation, 2014): Critical values are 1% (-5.34); 5% (-4.80) and 10% (-4.58)

Series (Level)	t-statistic	Optimal Breakpoints	Decisions
ΔlnOPEN	-7.928***	1993	Stationary
ΔlnGE	-6.176***	1984	Stationary
Δlny	-6.943***	1977	Stationary
$\Delta \ln M2$	-7.023***	1985	Stationary
ΔlnIN	-5.723***	1977	Stationary

Table 3 ZA (1992) unit root tests Results

(Author's computation, 2014); Critical values are 1% (-5.34); 5% (-4.80) and 10% (-4.58): Note: *** denotes statistical significance at the 10%, 5% and 1% levels respectively.

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4.2 Clemente-Montanes-Reyes Unit Root Test Results

In the current study the stationarity test was performed using the Clemente-Montanes-Reyes test (1998) that accounts for two structural breaks. The Clemente-Montanes-Reyes tests are the innovational outlier (IO) and Additive Outlier (AO). The IO model results are shown in Table 4 and that of AO model are exhibited in Table 5. According to the IO model, structural changes take place gradually and allows for a break in both the slope and the intercept of the model. However, in the AO test, the hypothesis is that structural changes are only in the slope and it is rapid. The test null assumption is that the variables are unit root with structural breaks. The alternative assumption is that the variables are not unit root with a break. Despite the structural break in the series using the innovative outlier, except, government expenditure, the null assumption is not rejected in levels

as shown in Table 4. In Table 5 except money supply, we can reject the null hypothesis of a unit root in the series variables in first difference and logarithm form.

IO mode: IO model, structural changes take place gradually and allows for a break in both the slope and the intercept of the model						
Series(levels)	Series(levels) t-statistic Optimal Breakpoints 5% critical value Decisions					
OPEN	-3.056	1984 & 1994	-5.49	Not stationary		
у	0.604	1999 & 2004	-5.49	Not stationary		
GE	-5.53	1979 & 1990	-5.49	Stationary		
M2	-2.184	1977 & 1990	-5.49	Not stationary		
IN	1.291	1998 & 2006	-5.49	Not stationary		

Table 4 Clemente et al., (1998) structural break with double mean shifts(Author's computation, 2014)

IO mode: IO model, structural changes take place gradually and allows for a break in both the slope and the intercept of the model					
Series	Series t-statistic Optimal Breakpoints 5% critical Decisions				
Δlny	-7.758	1975 & 1982	-5.49	Stationary	
ΔlnOPEN	-8.32	1981 & 1986	-5.49	Stationary	
ΔlnGE	-5.849	1981 & 2004	-5.49	Stationary	
$\Delta \ln M2$	-4.066	1975 & 1982	-5.49	Not stationary	
ΔlnIN	-6.617	1974 & 1982	-5.49	Stationary	

 Table 5 Clemente et al., (1998) structural break with double mean shifts (Author's computation, 2014)

In the case of the use of the additive outlier, in the face of the structural break in the variables in levels, the null assumption is not rejected in Table 6. In Table 7, except government expenditure, the null assumption is not rejected in first difference.

AO m	AO model: In the AO model, the assumption is that structural changes are rapid and allow for a break in only the slope.					
Series	Seriest-statisticOptimal Breakpoints5% critical valueDecisions					
GE	-3.568	1978 & 1988	-5.49	Unit root		
у	-1.234	2002 & 2007	-5.49	Unit root		
OPEN	-0.894	1988 & 1998	-5.49	Unit root		
M2	-3.623	1981 & 1996	-5.49	Unit root		
IN	-2.568	1998 & 2005	-5.49	Unit root		

 Table 6 Clemente et al., (1998) structural break with double mean shifts (Author's computation, 2014)

AO model: In	AO model: In the AO model, the assumption is that structural changes are rapid and					
Series (1 st dif.)	allow for a break in only the slope.Series (1st dif.)t-statisticOptimal Breakpoints5% cri. valueDecisions					
ΔlnGE	-5.56	1983 & 2003	-5.49	Stationary		
Δlny	-1.427	1974 & 1981	-5.49	Unit root		
ΔlnM2	-3.411	1977 & 1981	-5.49	Unit root		
ΔlnIN	0.177	1973& 1981	-5.49	Unit root		
ΔlnOPEN	-3.618	1980 & 2000	-5.49	Unit root		

 Table 7 Clemente et al., (1998) structural break with double mean shifts (Author's computation, 2014)

4.3 Results and Analysis of the Economic Growth Determinants

In this section of the paper, growth determinants results of the Gregory and Hansen (G-H) Co-integration approach is presented. The results on the estimated models [model C; model C/T; model C/S and model C/S/T] are reported in Table 8. The results revealed evidence of no significant cointegration in models C and model C/T but significant cointegration in model C/S and model C/S/T. Therefore, Models C/S/T and model C/S are estimated.

Regressors	Model	ADF	BP		
	С	-5.22	2001:32:00		
	C/T	-5.22	2001:32:00		
	C/S	-9.980***	2002:33:00		
y, M2, GE, OPEN, IN	C/S/T	-7.180***	1995:26:00		
	Note: The 1% CVs are -6.050 for ADF				
	Note: The 5% CVs are -5.560 for ADF				
	Note: The 10% CVs are -5.310 for ADF				

 Table 8 Gregory-Hansen Structural Break Cointegration Test (Economic Growth; Model-Model C; Model C/T; Model C/S and Model C/S/T)

Source: Author's computation, 2014. Note: ** denote significance at 5% level of significance

The appropriate model for the long-run estimates of the two models was estimated by employing the OLS test of regression. The results are presented in Table 9. The estimates of the two models seem to imply that Model C/S is the most plausible model since more estimated coefficients of the explanatory variables are significant. The short-run dynamic equation for the determinants of economic growth with the error-correction adjustment model (ECM) was estimated using the residuals from Model C/S. Growth increases by about 33.3%, about 35.9%, about 13.9% and about 0.06% (though insignificant) when money supply, government expenditure, trade openness and inflation increase by 1% respectively in the long run in model C/S.

Cointegration equations 1970-2011 (Economic Growth)				
Regressors	GH-C/S (2002)	GH-C/S/T (1995)		
Constant	12.886	18.745		
Constant	(19.940)***	(14.260)***		
Dum X Constant	0.017	0.016		
Duni A Constant	-0.816	-0.292		
Trend		0.071		
Tiena	n.a	(4.694)***		
Dum X Trend	n.a	n.a		
lnM2	0.333	-0.004		
1111112	(6.758)***	-0.053		
Dum X lnM2	n.a	n.a		
lnGE	0.359	0.088		
lige	(12.19)***	-1.462		
Dum X lnGE	n.a	n.a		
lnOPEN	0.138	0.088		
HOPEN	(5.576)***	(3.899)***		
Dum X lnOPEN	n.a	n.a		
lnIN	0.006	-0.167		
11111N	-0.839	(-4.435)***		
Dum X lnIN	n.a	n.a		
R^2	0.978	0.987		
Adjusted R-square	0.975	0.984		

 Table 9 Long Run Coefficients Estimates

 Source: author's computation, 2014. Note *** denote significance at 1% level

The short-run results of the coefficients estimated are reported in Table 10. All the independent variables are insignificant with the coefficient of inflation alone having unexpected a priori theoretical sign. The value of the error correction term (ECM) of -0.089 is statistically insignificant. However, it has the expected a priori theoretical negative of negative.

Model C/S	Iodel C/SRegresand = $\Delta \ln y$				
Regressors	Elasticities	Std Error	T-ratio	P-value	
Δlny-1	0.157	0.226	0.694	0.349	
ΔlnGE-1	0.043	0.065	0.659	0.514	
ΔlnOPEN-1	0.055	0.035	1.548	0.131	
$\Delta lnIN-1$	0.044	0.066	0.656	0.517	
$\Delta \ln M2$ -1	0.047	0.082	0.566	0.575	
ECM ₋₁	-0.057	0.059	0.97	0.339	
Constant	-1.02	1.073	-0.951	0.349	
Mean dependent var	0.033	S.D. dependent var	•	0.048	
Sum squared resid	0.068	S.E. of regression 0.04		0.045	
R-squared	0.241	Adjusted R-squared 0.102		0.102	
rho	0.078	Durbin-Watson		1.712	

Table 10 Short Run Coefficient Estimates Source: Author's computation, 2014

Figures 1 and Figure 2 depict the stability test results of the parameters, which show that the coefficients estimated are not stable according to the CUSUM test. The results of the CUSUMSQ test indicate the squared residuals are stable since the fall in 5% critical boundaries.



Figure 1 CUSUM



Figure 2 CUSUMSQ

5 DISCUSSIONS OF RESULTS

The paper examines and analyses the determinants of economic growth for Ghana for the period 1970 to 2011 using the Gregory and Hansen cointegration test which accounts for the effect of structural breaks.

On the nature of structural breaks in the data used, the study shows that there are structural breaks that coincide with identified climatic, economic, and political shocks. The findings are in support of that of Kiran, Yavus and Guris (2009) for Turkey; Binh (2011) for Vietnam; Dobnik (2011) for 23 OECD countries; and Dramani et al. (2012) for Ghana. The theoretical implications are that, the theory of structural breaks is supported and that forecasted values of real output that do not take account of structural breaks might have errors and are unreliable.

In the case of the determinants of economic growth, the findings of a positive effect of trade on growth are in line with that of previous research works (Manni & Afzal, 2012; Soliu & Ibrahim, 2014) that reported of a positive effect of trade on growth.

Besides, the findings of a positive nexus between financial development and growth are in line with the findings of earlier researchers (Mansor, 2005; Owoye et al., 2007; Mohammad et al., 2009; Nouri & Samimi, 2011; Tabi & Ondoa, 2011) who reported of a positive relationship between financial development and growth, but inconsistent with the findings of other researchers (Ahmed & Suliman, 2011) who reported of a negative effect of financial development and growth.

Further, the findings of a positive effect of inflation (but insignificant) on growth are in support of the works of previous researchers (Mallik & Chowdhury, 2001; Ahmed & Suliman, 2011) who reported of a positive effect of inflation on growth. However, the findings are not in support of previous studies (Fischer, 1993; Ghosh & Phillips, 1998; Shitundu & Luvanda, 2000; Mubarik, 2005; Bick, Kremer and Nautz, 2009; Mohammad et al., 2009; Sergii, 2009; Espinoza et al., 2010; Hasanov, 2011; Marbuah, 2010; Kasidi & Mwakanemela, 2013) that produced negative results. Other researchers (Frimpong & Oteng-Abayie, 2010) have reported of the neutral effect of inflation on economic growth, which is contrary to the findings of the current study.

Lastly, the empirical results on the positive effect of government expenditure on growth are consistent with the findings of previous researchers in the literature (Ranjan & Sharma, 2008; Alexiou, 2009; Ighodaro & Oriakhi, 2010; Adeniyi & Bashir, 2011; Srinivasan, 2013). According to

researchers, this might result from positive externalities through the harmonization of the conflicts between private and social interests and the provision of socially optimal direction for growth as well as offsetting market failures (Ghali, 1998). Ighodaro and Oriakhi (2010) and Adeniyi and Bashir (2011) reported a positive effect of government spending on economic growth in Nigeria. Srinivasan (2013) reported a statistically significant positive relationship between GDP and public expenditure in India.

The findings are not in line with the works of other researchers (Barro, 1991; Bajo-Rudio, 2000; Pevcin, 2003; Afonso & Furceri, 2008; Pham, 2009; Mohammad et al., 2009; Maku, 2009; Bergh & Karlsson, 2010; Afonso & Jalles, 2011) who have reported of a significant negative relationship between government expenditure and growth. According to these researchers, increasing government expenditure may deteriorate economic growth through the crowding-out effect. The private sector is crowded out as a result of distortions of the tax, government inefficiencies, incentives systems, and interventions to free markets system. The findings of the current paper are inconsistent with that of researchers (Taban, 2010; Verma & Arora, 2010) who reported an insignificant link between government expenditures and growth.

6. CONCLUSIONS AND RECOMMENDATIONS

The purpose of the study was to investigate the determinants of economic growth in the presence of structural breaks using the Gregory and Hansen cointegration model. The variables considered as the determinants were financial development, inflation, government expenditure, and trade openness.

The empirical findings are in agreement with the existence of cointegration in the presence of structural breaks. The study shows that there are structural breaks that coincide with identified climatic, economic, and political shocks. The theoretical implications are that the theory of structural breaks is supported and that forecasted values of real output that do not take account of structural breaks might have errors and are unreliable. In respect of policy, government-initiated structural reforms aimed at ensuring growth is of limited value, since the effect of such reforms on the long-run growth path will be offset by other shocks to the economy.

On the determinants of economic growth, the finding does not support short-run nexus between growth and the determinants considered in the study, though the conventional and new CUSUMSQ tests suggest the stability of equilibrium residuals which reinforces the cointegration nexus. This is so since there are no significant determinants of growth in the Gregory-Hansen model estimated in the short run. However, financial development, government expenditures, and trade openness are the long-run determinants of growth.

The importance of the current findings concerning policy formulation to achieve sustainable economic growth is that policymakers should put in place strategies to ensure that the financial sector is properly strengthening, trade is appropriately liberalised (*'open'*) and government expenditure is targeted at the productive sector of the economy. Future studies in line with the focus of the current study, based on panel cointegration, accounting for structural beaks effect, is worth embarking on.

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